

**ANALYTICAL PROOF OF ONE STRONG VERSION OF THE  
 LAW OF LARGE NUMBERS**

SH.K.FORMANOV AND B.B.KHUSAINOVA\*

ABSTRACT. An analytical proof based on the method of characteristic functions of Sakhonenko theorem on the validity of a strong option of the law of large numbers is presented in the paper: when a condition of the Lindeberg type is satisfied, the absolute value of the sum of independent random variables tends to zero on average. The mathematical expectation of the absolute value of a random variable uses the above proof through the corresponding characteristic function.

Let a "series scheme" of independent random variables (r.v.) be given in some probability space  $(\Omega, \mathfrak{F}, P)$  :

$$\{X_{n1}, \dots, X_{nn}\}, \quad n = 1, 2, \dots$$

and

$$S_n = X_{n1} + \dots + X_{nn}.$$

From the point of view of summation of independent r.v. results (available in theory), we can assume, without loss of generality, that

$$EX_{nj} = 0, \quad j = 1, 2, \dots, n. \tag{1}$$

Assume that

$$D_n(\alpha) = \sum_{j=1}^n E \min(|X_{nj}|, |X_{nj}|^{1+\alpha}), \quad \alpha > 0.$$

It is easy to check that if at certain  $\alpha = \alpha_0 > 0$

$$D_n(\alpha_0) \rightarrow 0, \quad n \rightarrow \infty,$$

then  $D_n(\alpha) \rightarrow 0$ , for all  $\alpha > 0$ .

Taking into account the latter, we assume that the following condition is met:

$$D_n = D_n(1) = \sum_{j=1}^n E \min(|X_{nj}|, |X_{nj}|^2) \rightarrow 0 \tag{D}$$

as  $n \rightarrow \infty$ .

In [1, p. 158] the following two versions of the law of large numbers (l.l.n.) are presented.

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The first version consists of a statement about the validity of the ordinary l.l.n. ( $S_n \xrightarrow{P} 0$ ) when conditions (1), (D) are met.

The second version, based on stronger convergence on average, is presented as the following theorem.

**Theorem A (A.I.Sakhonenko).** *When conditions (1), (D) are met,  $E|S_n| \rightarrow 0$  is true (or  $S_n \xrightarrow{P} 0$  which is the same). The statement of the above theorem means, obviously, the uniform integrability of the sequence  $\{S_n, n \geq 1\}$ , from it follows the ordinary version of the l.l.n.  $S_n \xrightarrow{P} 0$  so that*

$$P(|S_n| \geq \varepsilon) \leq \frac{E|S_n|}{\varepsilon} \rightarrow 0,$$

as  $n \rightarrow \infty$ .

The proof of Theorem A in [1, p. 159–160] was completed by the probabilistic method of "r.v.truncating". Here we give an "analytic proof" of Theorem A, based on the method of characteristic functions (ch.f.).

Assume that

$$\begin{aligned} F_{nj}(x) &= P(X_{nj} < x), \quad F_n(x) = P(S_n < x) \\ f_{nj}(t) &= Ee^{itX_{nj}}, \quad f_n(t) = Ee^{itS_n}, \quad j = 1, \dots, n. \end{aligned}$$

First we prove the following auxiliary statement.

**Lemma 0.1.** *For any r.v. with distribution function (d.f.)  $F(x)$*

$$E|X| = \int_{-\infty}^{\infty} |x| dF(x) = \frac{2}{\pi} \int_{-\infty}^{\infty} \frac{1 - \operatorname{Re} f(t)}{t^2} dt, \quad (2)$$

where  $f(t)$  is the ch.f. corresponding to the d.f.  $F(x)$ .

The assertion of the lemma is presented as an independent problem in [2, ch. 2, p. 409].

*Proof.* Changing the order of integration on the right-hand side of (2), we have

$$\begin{aligned} \int_{-\infty}^{\infty} \frac{1 - \operatorname{Re} f(t)}{t^2} dt &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} ((1 - \cos tx) dF(x)) t^{-2} dt = \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \left( \frac{1 - \cos tx}{t^2} dt \right) dF(x) = 2 \int_{-\infty}^{\infty} \left( \int_0^{\infty} \frac{1 - \cos tx}{t^2} dt \right) dF(x) \end{aligned} \quad (3)$$

Furthermore

$$\int_0^{\infty} \frac{1 - \cos tx}{t^2} dt = \int_0^T + \int_T^{\infty} = I_T^{(1)} + I_T^{(2)}. \quad (4)$$

Obviously, for any  $x \in R$

$$I_T^{(2)} = \int_T^{\infty} \frac{1 - \cos tx}{t^2} dt \leq 2 \int_T^{\infty} \frac{dt}{t^2} = \frac{2}{T}.$$

So

$$\sup_x I_T^{(2)}(x) \leq \frac{2}{T}. \quad (5)$$

After integrating by parts, we obtain:

$$I_T^{(1)}(x) = - \int_0^T (1 - \cos tx) d\left(\frac{1}{t}\right) = -\frac{1 - \cos Tx}{T} + x \int_0^T \frac{\sin tx}{t} dt. \quad (6)$$

Taking into account the last relations (5) and (6), we obtain:

$$\begin{aligned} \int_0^\infty \left( \int_0^\infty \frac{1 - \cos tx}{t^2} dt \right) dF(x) &= \int_0^\infty \lim_{T \rightarrow \infty} I_T^{(1)}(x) dF(x) = \\ &= \int_0^\infty x \left( \int_0^\infty \frac{\sin tx}{t} dt \right) dF(x) = \frac{\pi}{2} \int_0^\infty x dF(x). \end{aligned} \quad (7)$$

Here the well-known equality (the Dirichlet integral) is used

$$\int_0^\infty \frac{\sin tx}{t} dt = \begin{cases} \frac{\pi}{2}, & x > 0 \\ -\frac{\pi}{2}, & x < 0 \end{cases}, \quad (8)$$

Applying formulas (6) and (8) in a similar way, we obtain:

$$\int_{-\infty}^0 \left( \int_0^\infty \frac{1 - \cos tx}{t} dt \right) dF(x) = -\frac{\pi}{2} \int_{-\infty}^0 x dF(x) = \frac{\pi}{2} \int_{-\infty}^0 |x| dF(x) \quad (9)$$

Now from equalities (3), (7), (9) it follows that

$$\begin{aligned} \int_{-\infty}^\infty \frac{1 - \operatorname{Re} f(t)}{t^2} dt &= \int_{-\infty}^0 \left( \int_0^\infty \frac{1 - \cos tx}{t} dt \right) dF(x) + \int_0^\infty \left( \int_0^\infty \frac{1 - \cos tx}{t} dt \right) dF(x) = \\ &= \frac{\pi}{2} \int_{-\infty}^0 |x| dF(x) + \frac{\pi}{2} \int_0^\infty x dF(x) = \frac{\pi}{2} \int_{-\infty}^\infty |x| dF(x). \end{aligned}$$

Thus, formula (2) is proved.  $\square$

Now we will prove Theorem A using the lemma (or formula (2) which is the same) Assume that

$$E |S_n| = \frac{2}{\pi} \int_{-\infty}^\infty \frac{1 - \operatorname{Re} f_n(t)}{t^2} dt = \frac{2}{\pi} \left( I_n^{(1)} + I_n^{(2)} \right), \quad (10)$$

where

$$I_n^{(1)} = \int_{|t| \leq T} \frac{\operatorname{Re}(1 - f_n(t))}{t^2} dt, \quad I_n^{(2)} = \int_{|t| > T} \frac{\operatorname{Re}(1 - f_n(t))}{t^2} dt, \quad T > 0.$$

It is evident that

$$I_n^{(2)} = \int_{|t|>T} \left( \int_{-\infty}^{\infty} (1 - \cos tx) dF_n(x) \right) \frac{dt}{t^2} \leq 2 \int_{|t|>T} \frac{dt}{t^2} \leq \frac{8}{T}.$$

So

$$\limsup_{n \rightarrow \infty} I_n^{(2)} \leq \frac{8}{T}$$

and, by the randomness of  $T > 0$  we obtain

$$I_n^{(2)} \rightarrow 0, \quad n \rightarrow \infty. \quad (11)$$

Next, we use the following equality

$$\begin{aligned} 1 - f_n(t) &= \prod_{j=1}^n 1 - \prod_{j=1}^n f_{nj}(t) = \sum_{j=1}^n (1 - f_{nj}(t)) \prod_{k=j+1}^n f_{nk}(t) = \\ &= \sum_{j=1}^n (1 - f_{nj}(t)) \left[ 1 + \left( \prod_{k=j+1}^n f_{nk}(t) - \prod_{k=j+1}^n 1 \right) \right]. \end{aligned}$$

By virtue of this equality, we have

$$\frac{\operatorname{Re}(1 - f_n(t))}{t^2} = \sum_{j=1}^n \frac{\operatorname{Re}(1 - f_{nj}(t))}{t^2} + \psi_n(t), \quad (12)$$

where

$$\psi_n(t) = \sum_{j=1}^n \operatorname{Re} \left[ \frac{1 - f_{nj}(t)}{t^2} \left( \prod_{k=j+1}^n f_{nk}(t) - \prod_{k=j+1}^n 1 \right) \right].$$

By Lemma 3 from [1, ch.7, §6, p. 142]

$$\begin{aligned} \left| \prod_{k=j+1}^n f_{nk}(t) - \prod_{k=j+1}^n 1 \right| &\leq \sum_{k=j+1}^n |E(e^{itX_{nk}} - 1)| = \\ &= \sum_{k=j+1}^n |E(e^{itX_{nk}} - 1 - itX_{nk})| \leq 2h(t) \sum_{k=1}^n Eg(X_{nk}), \end{aligned}$$

where  $h(t) = \max(|t|, |t|^2)$ ,  $g(x) = \min(|x|, |x|^2)$ .

So, for each  $j = 1, \dots, n$

$$\begin{aligned} \left| \prod_{k=j+1}^n f_{nk}(t) - \prod_{k=j+1}^n 1 \right| &\leq 2h(t) \left[ \sum_{k=1}^n \left( \int_{|x| \leq 1} x^2 dF_{nk}(x) + \right. \right. \\ &\left. \left. + \int_{|x| > 1} |x| dF_{nk}(x) \right) \right] = 2h(t) (M_n + L_n) = 2h(t) D_n \rightarrow 0. \end{aligned} \quad (13)$$

Here

$$M_n = \sum_{k=1}^n \int_{|x| \leq 1} x^2 dF_{nk}(x), \quad L_n = \sum_{k=1}^n \int_{|x| > 1} |x| dF_{nk}(x).$$

Then, according to formula (6)

$$\begin{aligned} \sum_{j=1}^n \int_{-T}^T \frac{\operatorname{Re}(1 - f_{n_j}(t))}{t^2} dt &= -2 \sum_{j=1}^n \int_{-\infty}^{\infty} \left( \int_0^T (1 - \cos tx) d\left(\frac{1}{t}\right) \right) dF_{n_j}(x) \leq \\ &\leq 2 \left[ \sum_{j=1}^n \int_{-\infty}^{\infty} \left( \frac{1 - \cos Tx}{T} \right) dF_{n_j}(x) + \sum_{j=1}^n \int_{-\infty}^{\infty} x \left( \int_0^T \frac{\sin tx}{t} dt \right) dF_{n_j}(x) \right] \end{aligned} \quad (14)$$

It is easy to see that the following estimates hold

$$\begin{aligned} \sum_{j=1}^n \int_{-\infty}^{\infty} \left( \frac{1 - \cos Tx}{T} \right) dF_{n_j}(x) &= \sum_{j=1}^n \int_{|x| \leq 1} \left( \frac{1 - \cos Tx}{T} \right) dF_{n_j}(x) + \\ + \sum_{j=1}^n \int_{|x| > 1} \left( \frac{1 - \cos Tx}{T} \right) dF_{n_j}(x) &\leq T \sum_{j=1}^n \int_{|x| \leq 1} \frac{2 \sin^2 \frac{Tx}{2}}{4 \left( \frac{Tx}{2} \right)^2} \cdot x^2 dF_{n_j}(x) + \\ &+ \frac{2}{T} \sum_{j=1}^n \int_{|x| > 1} |x| dF_{n_j}(x) \leq C(T) (M_n + L_n) = C(T) D_n \rightarrow 0. \end{aligned} \quad (15)$$

Hereinafter,  $C(T)$  denotes a positive constant (depending on  $T$ ), which is different in different points.

Then

$$\begin{aligned} \sum_{j=1}^n \int_{-\infty}^{\infty} x \left( \int_0^T \frac{\sin tx}{t} dt \right) dF_{n_j}(x) &\leq \sum_{j=1}^n \int_{|x| \leq 1} |x| \left| \int_0^T \frac{\sin tx}{t} dt \right| dF_{n_j}(x) + \\ + \sum_{j=1}^n \int_{|x| > 1} |x| \left| \int_0^T \frac{\sin tx}{t} dt \right| dF_{n_j}(x) &\leq T \sum_{j=1}^n \int_{|x| \leq 1} |x|^2 dF_{n_j}(x) + \\ + \frac{\pi}{2} \sum_{j=1}^n \int_{|x| > 1} |x| dF_{n_j}(x) &\leq \max\left(T, \frac{\pi}{2}\right) (M_n + L_n) \leq C(T) D_n \rightarrow 0. \end{aligned} \quad (16)$$

Now from relations (14) - (16) we obtain:

$$\sum_{j=1}^n \int_{-T}^T \frac{\operatorname{Re}(1 - f_{n_j}(t))}{t^2} dt \leq C(T) D_n \rightarrow 0. \quad (17)$$

By relations (12), (13), and (17), we obtain

$$\int_{-T}^T \psi_n(t) dt \leq C(T) D_n^2 \rightarrow 0. \quad (18)$$

Finally, from relations (12), (17), and (18), we can conclude that

$$\int_{-T}^T \frac{\operatorname{Re}(1 - f_n(t))}{t^2} dt \leq C(T) (D_n + D_n^2).$$

Thus

$$I_n^{(1)} = \int_{-T}^T \frac{\operatorname{Re}(1 - f_n(t))}{t^2} dt = O(D_n + D_n^2) \rightarrow 0 \quad (19)$$

the proof of Theorem A follows from relations (11) and (19).

### References

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SH.K.FORMANOV: V.I.ROMANOVSKIY INSTITUTE OF MATHEMATICS, 81, MIRZO-ULUGBEK STR., TASHKENT, 100170, UZBEKISTAN

*E-mail address:* shakirformnov@yandex.ru

B.B.KHUSAINOVA: V.I.ROMANOVSKIY INSTITUTE OF MATHEMATICS, 81, MIRZO-ULUGBEK STR., TASHKENT, 100170, UZBEKISTAN

*E-mail address:* bikajon.khusainova@yandex.ru